

Math Recap

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Matrix Algebra and Geometric Constructs

$$Ax = \sum_i x_i a_i$$

$$\begin{bmatrix} | & & | \\ a_1 & \dots & a_k \\ | & & | \end{bmatrix} \begin{bmatrix} x_1 \\ \vdots \\ x_k \end{bmatrix}$$

algebraic

$$Bx = \begin{bmatrix} \overrightarrow{b_1^T} \\ \vdots \\ \overleftarrow{b_n^T} \end{bmatrix} = \begin{bmatrix} b_1^T x \\ \vdots \\ b_n^T x \end{bmatrix}$$

Geometric

Inner and Outer Products

Euclidean inner product: $\langle u, v \rangle = u^T v = \sum_i u_i \cdot v_i$

induces a norm: $\|u\|_2 = \sqrt{\langle u, u \rangle} = \sqrt{\sum_i u_i^2}$

Outer product: $uv^T = \begin{bmatrix} | & & | \\ u v_1 & \dots & u v_n \\ | & & | \end{bmatrix} \in \mathbb{R}^{m \times n}$ for $u \in \mathbb{R}^m, v \in \mathbb{R}^n$

Singular Value Decomposition

Theorem: Let $A \in \mathbb{R}^{m \times n}$. We can decompose A as a product

$A = U \Sigma V^T$, where $U \in \mathbb{R}^{m \times m}$ and $V \in \mathbb{R}^{n \times n}$ are orthogonal, and Σ is a $m \times n$ diagonal matrix with non-negative diagonal entries.

$$A = U \Sigma V^T$$

$$= \begin{bmatrix} u_1 & \dots & u_m \end{bmatrix} \begin{bmatrix} \sigma_1 & & \\ & \dots & \\ & & \sigma_m \end{bmatrix} \begin{bmatrix} -v_1^T \\ \vdots \\ -v_n^T \end{bmatrix}$$

$\sigma_i \geq 0$

$$\langle u_i, u_j \rangle = \begin{cases} 1 & , \text{if } i=j \\ 0 & , \text{otherwise.} \end{cases}$$

$$\langle v_i, v_j \rangle = \begin{cases} 1 & , \text{if } i=j \\ 0 & , \text{otherwise} \end{cases}$$

Singular Value Decomposition: Geometric Interpretation

$$Ax = U \Sigma V^T x$$



reflection/rotation

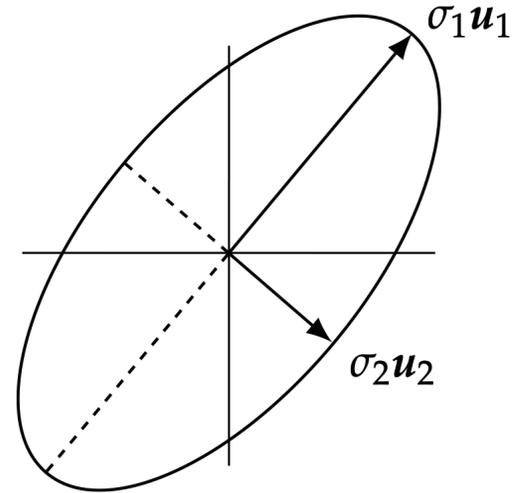
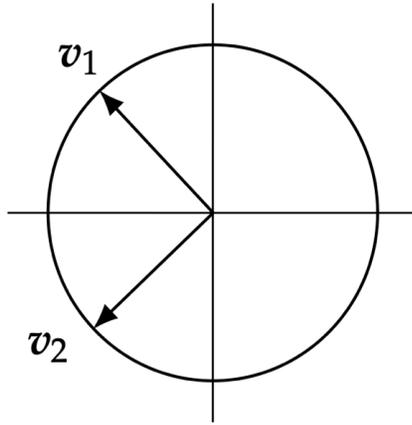


stretch/squeeze



rotation/reflection

Singular Value Decomposition: Illustration



Orthogonal Projections

Consider linearly independent vectors $a_1, \dots, a_n \in \mathbb{R}^m$ ($n \leq m$)

Q: What is the orthogonal projection y of a vector x onto the $\text{span}\{a_1, \dots, a_n\}$?

$$A = \begin{bmatrix} | & & | \\ a_1 & \dots & a_n \\ | & & | \end{bmatrix} \in \mathbb{R}^{m \times n}$$

$$\text{rank}(A) = n \quad (\text{full-rank})$$

Requirements

1) $y \in \text{span}\{a_1, \dots, a_n\} \Rightarrow y = Av$

2) $x - y \perp \text{span}\{a_1, \dots, a_n\} \Leftrightarrow \langle x - y, w \in \text{span}\{a_1, \dots, a_n\} \rangle = 0$

$$\Leftrightarrow \langle x - y, a_i \rangle = 0$$

$$\Leftrightarrow A^T(x - y) = 0$$

Orthogonal Projections: Continued

$$\begin{aligned} \text{From 2) } A^T(x-y) &= 0 \iff A^T x = A^T y \\ &= \underbrace{A^T A}_{\text{invertible since } a_1, \dots, a_n \text{ are linearly independent}} v \quad (\text{from 1}) \\ &\Rightarrow v = (A^T A)^{-1} A^T x \end{aligned}$$

$$\text{From 1) } y = Av = \underbrace{A(A^T A)^{-1} A^T}_{\text{projection matrix}} x =: \text{Proj}_A(x)$$

Eigenvalues, Eigenvectors

Let $\lambda \in \mathbb{R}$, $v \in \mathbb{R}^n$, $v \neq 0$. Let $A \in \mathbb{R}^{n \times n}$

$$Av = \lambda v$$

↑
eigenvalue

↙
eigenvector

Finding them: find roots of characteristic equation $\det(A - \lambda I) = 0$
If A is symmetric ($A^T = A$), then all eigenvalues are real.

Relationship to SVD: $\sigma_i^2(A) = \lambda_i(A^T A) = \lambda_i(A A^T)$

Spectral decomposition: $A = X \Lambda X^{-1}$

$$\begin{matrix} \nearrow & \nwarrow \\ \begin{bmatrix} | & & | \\ v_1 & \dots & v_n \\ | & & | \end{bmatrix} & \begin{bmatrix} \lambda_1 & & 0 \\ & \dots & \\ 0 & & \lambda_n \end{bmatrix} \end{matrix}$$

Quadratic Forms

Let $A \in \mathbb{R}^{n \times n}$ be symmetric

→ quadratic form: $x^T A x = \sum_i \sum_j a_{ij} \cdot x_i \cdot x_j$

A is positive definite (p.d.) if $x^T A x > 0 \quad \forall x \neq 0$

$\Leftrightarrow \lambda_i > 0 \quad \forall i=1, \dots, n$

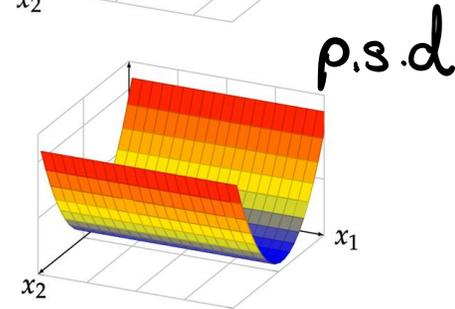
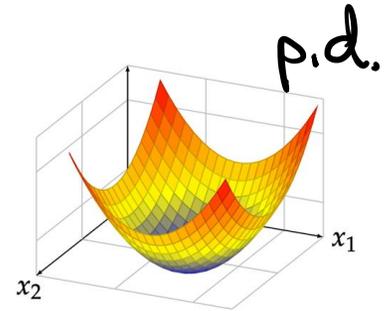
A is positive semi-definite (p.s.d.) if $x^T A x \geq 0 \quad \forall x \neq 0$

$\Leftrightarrow \lambda_i \geq 0 \quad \forall i=1, \dots, n$

Cholesky decomposition (square root): If A p.s.d., then A can be factored as

$A = L L^T$, where L is a lower-triangular matrix. $L = \begin{bmatrix} \blacksquare & & 0 \\ & \blacksquare & \\ & & \blacksquare \end{bmatrix}$

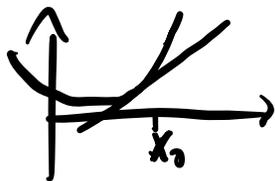
$$x \mapsto x^T A x$$



Multivariate Derivatives

$$f: \mathbb{R} \rightarrow \mathbb{R}$$

$$f: \mathbb{R}^n \rightarrow \mathbb{R}^m$$



What is the best linear approximation?

$$f(x_0 + \varepsilon) \approx f(x_0) + \underbrace{Df(x_0)[\varepsilon]}_{\substack{\text{linear map / matrix} \\ \text{matrix-vector product}}} \leftarrow \text{"Jacobian"}$$

$$[Df(x_0)]_{ij} = \frac{\partial f_i}{\partial x_j}(x_0) \quad \text{If } m=1, \text{ then } (Df(x_0))^T = \nabla f(x_0) \text{ "gradient"}$$

$$Df(x_0) \in \mathbb{R}^{m \times n} \quad (\mathbb{R}^n \rightarrow \mathbb{R}^m)$$

$$\nabla f(x_0) \in \mathbb{R}^n$$

$$m=1: \begin{array}{l} \text{derivative: } 1 \boxed{} \\ \text{gradient: } \boxed{}^1 \end{array}$$

Chain Rule

$$f: \mathbb{R}^n \rightarrow \mathbb{R}^m, \quad g: \mathbb{R}^k \rightarrow \mathbb{R}^n$$

$$D(f \circ g)(x_0) = Df(\underbrace{g(x_0)}_{\text{const}}) \circ Dg(x_0) \stackrel{\text{linear map}}{=} Df(g(x_0)) Dg(x_0)$$

Example:

$$\begin{aligned} \frac{d}{dt} f(x_0 + t \cdot v) \Big|_{t=0} &= D(f \circ g)(0) = Df(\underbrace{g(0)}_{x_0}) Dg(0) \\ &= \left[\frac{\partial f}{\partial x_1}(x_0) \quad \dots \quad \frac{\partial f}{\partial x_n}(x_0) \right] v \\ &= (\nabla f(x_0))^T v = \langle \nabla f(x_0), v \rangle \end{aligned}$$

Annotations in the original image:
- "vectors" with arrows pointing to x_0 and v
- "scalar" with an arrow pointing to t
- $x_0 \in \mathbb{R}^{1 \times n}$ (green)
- $v \in \mathbb{R}^n$ (orange)

Taylor Expansion and Second-order Derivatives

$$f: \mathbb{R}^n \rightarrow \mathbb{R}$$

first-order expansion at x_0 :

$$f(x) = f(x_0) + \langle \nabla f(x_0), x - x_0 \rangle + o(\|x - x_0\|)$$

lecture notes

second-order expansion at x_0 :

$$f(x) = f(x_0) + \langle \nabla f(x_0), x - x_0 \rangle + \frac{1}{2} (x - x_0)^T D^2 f(x_0) (x - x_0) + o(\|x - x_0\|^2)$$

$$\underbrace{D^2 f(x_0)}_{\text{matrix}} = \begin{bmatrix} \frac{\partial^2 f}{\partial x_1^2}(x_0) & \dots & \frac{\partial^2 f}{\partial x_n \partial x_1}(x_0) \\ \vdots & \ddots & \vdots \\ \frac{\partial^2 f}{\partial x_1 \partial x_n}(x_0) & \dots & \frac{\partial^2 f}{\partial x_n^2}(x_0) \end{bmatrix}$$

"Hessian"

Random Variables, Distributions and Densities

Probability space $(\Omega, \mathcal{F}, IP)$ probability function: $IP: \mathcal{F} \rightarrow [0, 1]$

sample space
(atomic events) \uparrow
event space
 $\mathcal{F} \subseteq 2^\Omega$

Kolmogorov axioms:
1) $0 \leq IP(A) \leq 1 \quad \forall A \in \mathcal{F}$

2) $IP(\Omega) = 1$

3) For a countable mutually disjoint set of events

$\{A_i \in \mathcal{F}\}_i \quad (A_i \cap A_j = \emptyset)$

$$IP\left(\bigcup_{i=1}^{\infty} A_i\right) = \sum_{i=1}^{\infty} IP(A_i)$$

Random variable: $X: \Omega \rightarrow \mathbb{R}$

\hookrightarrow deterministic for given input $\omega \in \Omega$

distribution: $IP_X(A) = IP(\{\omega \in \Omega : X(\omega) \in A\})$

\hookrightarrow subset of values $A \subseteq \mathbb{R}$

cumulative distribution function (CDF) $F_X(x) = IP((-\infty, x]) = IP(X \leq x)$

probability density function (PDF)

$p_X: \mathbb{R} \rightarrow [0, \infty)$, such that for any $I \subseteq \mathbb{R}$

$$IP(I) = \int_I p_X(x) dx$$

Joint Distributions

Let X, Y be two random variables on the same $(\Omega, \mathcal{F}, \mathbb{P})$

joint distribution: $\mathbb{P}_{X,Y}(A) = \mathbb{P}(\{\omega \in \Omega : (X(\omega), Y(\omega)) \in A\})$, for $A \in \mathbb{R}^2$

joint CDF: $F_{X,Y}(x,y) = \mathbb{P}_{X,Y}((-\infty, x] \times (-\infty, y]) = \mathbb{P}(X \leq x, Y \leq y)$

joint PDF: $p_{X,Y}(x,y)$, such that $\mathbb{P}_{X,Y}(A) = \int_A p_{X,Y}(x,y) dx dy$

Marginal Distributions

Given joint distribution $P_{X,Y}$, what is the distribution P_X ?

"marginal" distribution: $P_X(A) = P_{X,Y}(A \times \mathbb{R})$ for $A \subseteq \mathbb{R}$

given the joint PDF $p_{X,Y}$, marginal PDF p_X :

$$p_X(x) = \int_{\mathbb{R}} p_{X,Y}(x,y) dy$$

Conditional Distributions and Independence

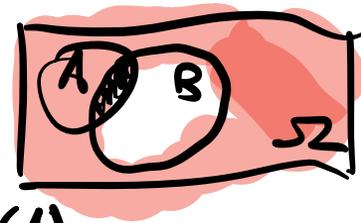
Let's consider two events A, B .

events A and B are independent iff $IP(A \cap B) = IP(A) \cdot IP(B)$

$$p_{X,Y}(x,y) = p_X(x) \cdot p_Y(y)$$

if $IP(B) > 0$, the conditional probability of A given B is

$$IP(A|B) = \frac{IP(A \cap B)}{IP(B)}$$



Bayes' rule: $IP(A|B) = \frac{IP(A \cap B)}{IP(B)} = \frac{IP(B|A) \cdot IP(A)}{IP(B)}$

Expected Value

Expectation: $E[X] = \int_{\mathbb{R}} x \cdot p_X(x) dx$

Variance: $\text{Var}(X) = E[(X - E[X])^2] = E[X^2] - E[X]^2$

If X, Y independent, then $E[X \cdot Y] = E[X] \cdot E[Y]$

Conditional expectation $E[X | Y=y] = \int_{\mathbb{R}} x \cdot p_{X|Y}(x|y) dx$