



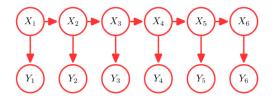
Review Session - PAI 2018 - Temporal Models

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Learning & Adaptive Systems group Institute of Machine Learning ETH Zürich

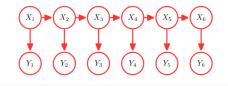
Hidden Markov models

- Instead of assuming all variables are observed (*Y*_t), we assume we have variables *X*_t, which we cannot observe.
- Assume: A graphical model for Hidden Markov Model (HMM)



- $\{X_t\}_t$ are called states and $\{Y_t\}_t$ are called observations.
- In this course: distributions assumed to be $P(X_t|X_{t+1}), P(Y_t|X_t)$ categorical or Gaussian mostly, however can be arbitrary.

What Inference Questions can we ask?



• Bayesian Filtering

 $P(X_t|Y_{1:t})$

• Prediction

 $P(X_{t+\tau}|Y_{1:t})$ where $\tau \geq 1$

• Smoothing (similar to filtering but on past data)

 $P(X_{\tau}|Y_{1:t})$ where $\tau < t$

• Most Probable Explanation (MPE)

Review Session - PAI 2018 - Temporal Mødels = arg max $P(\hat{X}_{1,i}|Y_{1,i})$

- Prior $P(X_1)$ needed.
- Bayesian Filtering usually done recursively. Assume you have P(X_t|Y_{1:t-1}), get P(X_t|Y_{1:t}) using Markov property and observation probabilities.
- Prediction (for one step; can be generalized)

$$P(X_{t+1}|Y_{1:t}) = \sum_{x} P(X_{t+1}, X_t = x|Y_{1:t}) = \sum_{x} P(X_{t+1}|X_t = x) P(X_t|Y_{1:t})$$

- Smoothing can be recast as calculating a marginal! HMM being a polytree, we know that Belief propagation is applicable and converges fast with forward-backward passes!
- MPE Sum product algorithm (search largest product).

- A special case of HMM is Kalman filter, when X_t and Y_t are Gaussian with Linear dynamics and observation model.
- Motion model = System dynamics $P(X_{t+1}|X_t)$ and is modeled using,

$$X_{t+1} = \mathbf{F}X_t + \epsilon_t$$

where, $\epsilon_t \sim \mathcal{N}(0, \Sigma_x)$. N.B.: Sum of Normal RV is a Normal RV!

• Sensor model = Observations $P(Y_t|X_t)$,

$$Y_t = \mathbf{H}X_t + \nu_t$$

where, $\nu_t \sim \mathcal{N}(0, \Sigma_x)$. N.B.: ϵ_t and ν_t are multivariate normals!

Kalman Filter II - How to?

- In general, this is the same as previously; but now the domain is different. Previously categorical distributions now continuous.
- Filtering:

$$P(X_t|Y_t = y_{1:t}) = \frac{1}{Z}P(X_t|Y_{t-1} = y_{1:t-1})P(Y_t = y_t|X_t)$$

• Prediction:

$$P(X_{t+1}|Y_{1:t} = y_{1:t}) = \int P(X_{t+1}|X_t = x)P(X_t = x|Y_{1:t} = y_{1:t})dx$$

- All these are Gaussian integrals; that can be solved exactly.
- We can calculate the Kalman filter uncertainty offline. Independent of y_{1:t}.
 - Review my multivariate Gaussians!

$$\mu_{A|B} = \mu_A + \Sigma_{AB} \Sigma_{BB}^{-1} (\mathbf{x}_B - \mu_B)$$

$$\Sigma_{A|B} = \Sigma_{AA} - \Sigma_{AB} \Sigma_{BB}^{-1} \Sigma_{BA}$$

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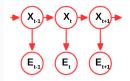
- What if the "motion" model is nonlinear? For example Particle filters.
- Idea: Pick samples $\{x_i\}$ that are propagated through the dynamics and then produce a histogram! Samples: $\{x_{t,i}\}_{i=1}^N$ from distribution $P(X_t|y_{1:t}) \approx \frac{1}{N} \sum_{i=1}^N \delta_{x,x_{t,i}}$.
- Propagate particles $x'_i \sim P(X_{t+1}|X_t, x_{t,i})$, calculate weights $w_i = \frac{1}{Z}P(y_{t+1}|x'_i)$ (transition model known).
- Next step $x_{i,t+1} \sim \frac{1}{N} \sum_{i=1}^{N} w_i \delta_{x,x_{t,i}}$
- Why this re-weighting? If we do not do it; we have a mode collapse (out of scope of the course).

5 [12 points] Temporal Models: Exploratory Rover

An exploratory rover is navigating through planet Vulcan, a deserted planet scattered with volcanic vents and radioactive valleys. The rover is equipped with a thermometer that registers only two levels, *hot* and *cold*. The rover sends back thermal responses E = hot when it is at a Volcanic vent (V), or a Radioactive valley (R), and E = cold when it is at Normal area (N). There is no chance of a mistaken reading.

The rover can only stay in one area on any given day. It travels around according to the following transition probabilities:

	$P(X_t \mid X_{t-1} = N)$	$P(X_t \mid X_{t-1} = V)$	$P(X_t \mid X_{t-1} = R)$
$X_t = N$	0.7	0.6	0.2
$X_t = V$	0.2	0.3	0.2
$X_t = R$	0.1	0.1	0.6



1. Define
$$\mathbf{X}_t = (P(X_t = N), P(X_t = V), P(X_t = R))$$

 $\mathbf{T} = \begin{pmatrix} 0.7 & 0.6 & 0.2 \\ 0.2 & 0.3 & 0.2 \\ 0.1 & 0.1 & 0.6 \end{pmatrix}$, then $\mathbf{X}_t \mathbf{T} = \mathbf{X}_{t+1}$.

- 2. Further, we know that observation model for Y as well. $P(E_t = hot|X_t = V) = P(E_t = hot|X_t = R) = 1$ and $P(E_t = cold|X_t = N) = 1$. So very simple model.
- 3. a) Observe {cold, hot hot}, what is $P(X_{1:3} = (NNN)|E_{1:3} = (CHH))$?

Using d-separation rules $P(X_{1:3}|E_{1:3}) = \underbrace{P(X_3|X_2, E_3)}_{P(E_3|X_3)P(X_3|X_2)/P(X_2, E_3)} P(X_2|X_1, E_{2:3})P(X_1|E_{1:3})$ $P(E_3 = hot|X_3 = N) = 0, \text{ so the overall probability is zero.}$ b) Given E{cold,hot,hot,cold} what is the most likely {X_{1:4}}. Hint: Not many calculations.

This is a MPE query, in other words,

$$X_{1:4} = \arg \max_{\hat{X}_{1:4}} P(\hat{X}_{1:4}|Y_{1:4})$$

Given that in cold state can be only observed when $X_i = N$, means that $X_1 = N$, $X_4 = N$. Now, we need to look which state is more probable for the middle points: VV, VR, RV, RR. We can calculate these probabilities using the transition matrix **T** i.e. $P(X_2 = V, X_3 = R, X_1 = N, X_4 = N) = P(X_2 = V|X_1 = N)P(X_3 = R|X_2 = V)P(X_4 = N|X_3 = R)$

Exam 2013 Problem 5 - IV

- c) We use particle filter with N = 8 particles.
 - 5 N (*i* ∈ 1,...5)
 - 2 V (*i* ∈ 6,7)
 - 1 R (*i* = 8)

The rover sends back E = hot. What are the weights for the next sampling?

From the previous slides, we know

$$w_i \propto P(E = \text{hot}|X = x_i)$$

For example for $w_1 \propto P(E = hot | X = N) = 0$. Similar for others. Weights need to sum to 1 so we can normalize them.