
Stochastic Bandits with Context Distributions

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Abstract

We introduce a novel stochastic contextual bandit model, where at each step the adversary chooses a distribution over a context set. The learner observes only the context distribution while the exact context realization remains hidden. This allows for a broader range of applications, for instance when the context itself is based on predictions. By leveraging the UCB algorithm to this setting, we propose an algorithm that achieves a $\tilde{O}(d\sqrt{T})$ high-probability regret bound for linearly parametrized reward functions. Our results strictly generalize previous work in the sense that both our model and the algorithm reduce to the standard setting when the environment chooses only Dirac delta distributions and therefore provides the exact context to the learner. We further obtain similar results for a variant where the learner observes the realized context after choosing the action, and we extend the results to the kernelized setting. Finally, we demonstrate the proposed method on synthetic and real-world datasets.

1 Introduction

In the contextual bandit model a learner interacts with an environment in several rounds. At the beginning of each round, the environment provides a context, and in turn, the learner chooses an action which leads to an a priori unknown reward. The learner's goal is to choose actions that maximize the cumulative reward, and eventually compete with the best mapping from context observations to actions. This model creates a dilemma of *exploration and exploitation*, as the learner needs to balance exploratory actions that provide information about the environment's reward structure, and exploitative actions that maximize the total return. Contextual bandit algorithms have been successfully used in many applications, including online advertisement, recommender systems and experimental design.

The contextual bandit model, as usually considered in the literature, assumes that the context is observed *exactly*. This is not always the case in applications, for instance, when the context is itself a noisy measurement or a forecasting mechanism. An example of such a context could be a weather or stock market prediction. In other cases such as recommender systems, privacy constraints can restrict access to certain user features, but instead we might be able to infer a distribution over those. To allow for uncertainty in the context, we consider a setting where the environment provides a *distribution over the context set*. The exact context is assumed to be a sample from this distribution, but remains hidden from the learner. Such a model, to the best of our knowledge, has not been considered in the literature before. Not knowing the context realization makes the learning problem more difficult, because the learner needs to estimate the reward function from noisy observations and without knowing the exact context that generated the reward. Our setting recovers the classical contextual bandit setting when the context distribution is a Dirac delta distribution. We also consider a natural variant of the problem, where the exact context is observed *after* the player has chosen the action. This allows for different applications, where at the time of decision the context needs to be predicted (e.g. weather conditions), but when the reward is obtained, also the exact context can be measured.

We focus on the setting where the reward function is linear in terms of action-context feature vectors. For this case, we show that by leveraging the UCB algorithm on a specifically designed bandit instance *without* feature uncertainty recovers an $\mathcal{O}(d\sqrt{T})$ high-probability regret bound. Our analysis includes a practical variant of the algorithm that requires only sampling access to the context distributions provided by the environment. We also extend our results to the kernelized setting, where the reward function can be any member of a known reproducing kernel Hilbert space (RKHS). For this case, we highlight an interesting connection to distributional risk minimization and we show that the natural estimator for the reward function is based on so-called kernel mean embeddings.

2 Stochastic Bandits with Context Distributions

We formally define the setting of *stochastic bandits with context distributions* that we outlined in the introduction. Let \mathcal{X} be a set of actions and \mathcal{C} a context set. The environment is defined by a fixed, but unknown reward function $f : \mathcal{X} \times \mathcal{C} \rightarrow \mathbb{R}$. At iteration $t \in \mathbb{N}$, the environment chooses a distribution $\mu_t \in \mathcal{P}(\mathcal{C})$ over the context set and samples a context realization $c_t \sim \mu_t$. The learner observes only μ_t but not c_t , and then chooses an action $x_t \in \mathcal{X}$. We allow that an *adaptive adversary* chooses the context distribution, that is μ_t may in an arbitrary way depend on previous choices of the learner up to time t . Given the learner's choice x_t , the environment provides a reward $y_t = f(x_t, c_t) + \epsilon_t$, where ϵ_t is σ -subgaussian, additive noise. The learner's goal is to maximize the cumulative reward $\sum_{t=1}^T f(x_t, c_t)$, or equivalently, minimize the cumulative regret

$$\mathcal{R}_T = \sum_{t=1}^T f(x_t^*, c_t) - f(x_t, c_t) \quad (1)$$

where $x_t^* = \arg \max_{x \in \mathcal{X}} \mathbb{E}_{c \sim \mu_t} [f(x, c)]$ is the best action provided that we know f and μ_t , but not c_t . Note that this way, we compete with the best possible mapping $\pi^* : \mathcal{P}(\mathcal{C}) \rightarrow \mathcal{X}$ from the observed context distribution to actions, that maximizes the expected reward $\sum_{t=1}^T \mathbb{E}_{c_t \sim \mu_t} [f(\pi^*(\mu_t), c_t) | \mathcal{F}_{t-1}, \mu_t]$ where $\mathcal{F}_t = \{(x_s, \mu_s, y_s)\}_{s=1}^t$ is the filtration that contains all information available at the end of round t . It is natural to ask if it is possible to compete with the stronger baseline that chooses actions given the context realization c_t , i.e. $\tilde{x}_t^* = \arg \max_{x \in \mathcal{X}} f(x, c_t)$. While this can be possible in special cases, a simple example shows, that in general the learner would suffer $\Omega(T)$ regret. In particular, assume that $c_t \sim \text{Bernoulli}(0.5)$, and $\mathcal{X} = \{0, 1\}$. Let $f(0, c) = c$ and $f(1, c) = 1 - c$. Clearly, any policy that does not know the realizations c_t , must have $\Omega(T)$ regret when competing against \tilde{x}_t^* .

From now on, we focus on linearly parameterized reward functions $f(x, c) = \phi_{x,c}^\top \theta$ with given feature vectors $\phi_{x,c} \in \mathbb{R}^d$ for $x \in \mathcal{X}$ and $c \in \mathcal{C}$, and a fixed parameter vector $\theta \in \mathbb{R}^d$. This setup is commonly referred to as the *linear bandit* setting. For the analysis we require standard boundedness assumptions $\|\phi_{x,c}\|_2 \leq 1$ and $\|\theta\|_2 \leq 1$ that we set to 1 for the sake of simplicity. In Section 4.2, we further consider a variant of the problem, where the learner observes c_t *after* taking the decision x_t . This simplifies the estimation problem, because we have data $\{(x_t, c_t, y_t)\}$ with exact context c_t available, just like in the standard setting. The exploration problem however remains subtle as at the time of decision the learner still knows only μ_t and not c_t . In Section 4.3 we extend our algorithm and analysis for *kernelized bandits* where $f \in \mathcal{H}$ is a member of a reproducing kernel Hilbert space \mathcal{H} . Related work is discussed in Section 6.

3 Background

We briefly review standard results from the linear contextual bandit literature and the upper confidence bound (UCB) algorithm that we built on later (Abbasi-Yadkori et al., 2011). The *linear contextual bandit* setting can be defined as a special case of our setup, where the choice of μ_t is restricted to Dirac delta distributions $\mu_t = \delta_{c_t}$, and therefore the learner knows beforehand the exact context which is used to generate the reward. In an equivalent formulation, the environment provides at time t a set of action-context feature vectors $\Psi_t = \{\phi_{x,c_t} : x \in \mathcal{X}\} \subset \mathbb{R}^d$ and the algorithm chooses an action x_t with corresponding features $\phi_t := \phi_{x_t, c_t} \in \Psi_t$. We emphasize that in this formulation the context c_t is extraneous to the algorithm, and everything can be defined in terms of the time-varying action-feature sets Ψ_t . As before, the learner obtains a noisy reward observation $y_t = \phi_t^\top \theta + \epsilon_t$ where ϵ_t is conditionally ρ -subgaussian with *variance proxy* ρ , i.e.

$$\forall \lambda \in \mathbb{R}, \quad \mathbb{E}[e^{\lambda \epsilon_t} | \mathcal{F}_{t-1}, \phi_t] \leq \exp(\lambda^2 \rho^2 / 2).$$

Algorithm 1 UCB for linear stochastic bandits with context distributions

Initialize $\hat{\theta} = 0 \in \mathbb{R}^d$, $V_0 = \lambda \mathbf{I} \in \mathbb{R}^{d \times d}$

For step $t = 1, 2, \dots, T$:

 Environment chooses $\mu_t \in \mathcal{P}(\mathcal{C})$ // context distribution
 Learner observes μ_t

 Set $\Psi_t = \{\bar{\psi}_{x, \mu_t} : x \in \mathcal{X}\}$ with $\bar{\psi}_{x, \mu_t} := \mathbb{E}_{c \sim \mu_t}[\phi_{x, c}]$ // variant 1, expected version
 Alternatively, sample $c_{t,1}, \dots, c_{t,L}$ for $L = t$, // variant 2, sampled version
 Set $\Psi_t = \{\tilde{\psi}_{x, \mu_t} : x \in \mathcal{X}\}$ with $\tilde{\psi}_{x, \mu_t} = \frac{1}{L} \sum_{l=1}^L \phi_{x, c_l}$

 Run UCB step with Ψ_t as context set // reduction

 Choose action $x_t = \arg \max_{\psi_{x, \mu_t} \in \Psi_t} \psi_{x, \mu_t}^\top \hat{\theta}_{t-1} + \beta_t \|\psi_{x, \mu_t}\|_{V_{t-1}^{-1}}$ // UCB action

 Environment provides $y_t = \phi_{x_t, c_t}^\top \theta + \epsilon$ where $c_t \sim \mu_t$ // reward observation

 Update $V_t = V_{t-1} + \psi_{x_s, \mu_s} \psi_{x_s, \mu_s}^\top$, $\hat{\theta}_t = V_t^{-1} \sum_{s=1}^t \psi_{x_s, \mu_s} y_s$ // least-squares update

Also here, the standard objective is to minimize the cumulative regret $\mathcal{R}_T = \sum_{t=1}^T \phi_t^{*\top} \theta - \phi_t^\top \hat{\theta}_t$ where $\phi_t^* = \arg \max_{\phi \in \Psi_t} \phi^\top \theta$ is the feature vector of the best action at time t .

To define the UCB algorithm, we make use of the confidence sets by Abbasi-Yadkori et al. (2011) for online least square regression. At the end of round t , the algorithm has adaptively collected data $\{(\phi_1, y_1), \dots, (\phi_t, y_t)\}$ that we use to compute the regularized least squares estimate $\hat{\theta}_t = \arg \min_{\theta' \in \mathbb{R}^d} \sum_{s=1}^t (\phi_s^\top \theta' - y_s)^2 + \lambda \|\theta'\|_2^2$ with $\lambda > 0$. We denote the closed form solution by $\hat{\theta}_t = V_t^{-1} \sum_{s=1}^t \phi_s y_s$ with $V_t = V_{t-1} + \phi_t \phi_t^\top$, $V_0 = \lambda \mathbf{I}_d$ and $\mathbf{I}_d \in \mathbb{R}^{d \times d}$ is the identity matrix.

Lemma 1 (Abbasi-Yadkori et al. (2011)). *For any stochastic sequence $\{(\phi_t, y_t)\}_t$ and estimator $\hat{\theta}_t$ as defined above for ρ -subgaussian observations, with probability at least $1 - \delta$, at any time $t \in \mathbb{N}$,*

$$\|\theta - \hat{\theta}_t\|_{V_t} \leq \beta_t \quad \text{where } \beta_t = \beta_t(\rho, \delta) = \rho \sqrt{2 \log \left(\frac{\det(V_t)^{1/2}}{\delta \det(V_0)^{1/2}} \right)} + \lambda^{1/2} \|\theta\|_2.$$

Note that the size of the confidence set depends the variance proxy ρ , which will be important in the following. In each round $t + 1$, the UCB algorithm chooses an action ϕ_{t+1} , that maximizes an upper confidence bound on the reward,

$$\phi_{t+1} := \arg \max_{\phi \in \Psi_{t+1}} \phi^\top \hat{\theta}_t + \beta_t \|\phi\|_{V_t^{-1}}.$$

The following result shows that the UCB policy achieves sublinear regret (Dani et al., 2008; Abbasi-Yadkori et al., 2011).

Lemma 2. *In the standard contextual bandit setting with ρ -subgaussian observation noise, the regret of the UCB policy with $\beta_t = \beta_t(\rho, \delta)$ is bounded with probability $1 - \delta$ by*

$$\mathcal{R}_T^{UCB} \leq \beta_T \sqrt{8T \log \left(\frac{\det V_T}{\det V_0} \right)}$$

The data-dependent terms can be further upper-bounded to obtain $\mathcal{R}_T^{UCB} \leq \tilde{\mathcal{O}}(d\sqrt{T})$ up to logarithmic factors in T (Abbasi-Yadkori et al., 2011, Theorem 3). A matching lower bound is given by Dani et al. (2008, Theorem 3).

4 UCB with Context Distributions

In our setting, where we only observe a context distribution μ_t (e.g. a weather prediction) instead of the context c_t (e.g. realized weather conditions), also the features ϕ_{x, c_t} (e.g. the last layer of a neural network that models the reward $f(x, c) = \phi_{x, c_t}^\top \theta$) are uncertain. We propose an approach that

transforms the problem such that we can directly use a contextual bandit algorithm as for the standard setting. Given the distribution μ_t , we define a new set of feature vectors $\Psi_t = \{\bar{\psi}_{x,\mu_t} : x \in \mathcal{X}\}$, where we denote by $\bar{\psi}_{x,\mu_t} = \mathbb{E}_{c \sim \mu_t}[\phi_{x,c} | \mathcal{F}_{t-1}, \mu_t]$ the expected feature vector of action x under μ_t . Each feature $\bar{\psi}_{x,\mu_t}$ corresponds to exactly one action $x \in \mathcal{X}$, so we can use Ψ_t as feature context set at time t and use the UCB algorithm to choose an action x_t . The choice of the UCB algorithm here is only for the sake of the analysis, but any other algorithm that works in the linear contextual bandit setting can be used. The complete algorithm is summarized in Algorithm 1. We compute the UCB action x_t with corresponding expected features $\psi_t := \bar{\psi}_{x_t,t} \in \Psi_t$, and the learner provides x_t to the environment. We then proceed and use the reward observation y_t to update the least squares estimate. That this is possible is not immediate, because y_t is a noisy observation of $\phi_{x_t,c_t}^\top \theta$, whereas UCB expects the reward $\psi_t^\top \theta$. We address this issue by constructing the feature set Ψ_t in such a way, that y_t acts as unbiased observation also for the action choice ψ_t . As computing exact expectations can be difficult and in applications often only sampling access of μ_t is possible, we also analyze a variant of Algorithm 1 where we use finite sample averages $\tilde{\psi}_{x,\mu_t} = \frac{1}{L} \sum_{l=1}^L \phi_{x,\tilde{c}_l}$ for $L \in \mathbb{N}$ i.i.d. samples $\tilde{c}_l \sim \mu_t$ instead of the expected features $\bar{\psi}_{x,\mu}$. The corresponding feature set is $\tilde{\Psi}_t = \{\tilde{\psi}_{x,\mu_t} : x \in \mathcal{X}\}$. For both variants of the algorithm we show the following regret bound.

Theorem 1. *The regret of Algorithm 1 with expected feature set Ψ_t and $\beta_t = \beta_t(\sqrt{4 + \sigma^2}, \delta/2)$ is bounded at time T with probability at least $1 - \delta$ by*

$$\mathcal{R}_T \leq \beta_T \sqrt{8T \log \left(\frac{\det V_T}{\det V_0} \right)} + 4\sqrt{2T \log \frac{4}{\delta}}.$$

Further, for finite action sets \mathcal{X} , if the algorithm uses sampled feature sets $\tilde{\Psi}_t$ with $L = t$ and $\beta_t = \tilde{\beta}_t$ as defined in (13), Appendix A.2, then with probability at least $1 - \delta$,

$$\mathcal{R}_T \leq \tilde{\beta}_T \sqrt{8T \log \left(\frac{\det V_T}{\det V_0} \right)} + 4\sqrt{2T \log \frac{2|\mathcal{X}|\pi T}{3\delta}}.$$

As before, one can further upper bound the data-dependent terms to obtain an overall regret bound of order $\mathcal{R}_T \leq \tilde{\mathcal{O}}(d\sqrt{T})$, see (Abbasi-Yadkori et al., 2011, Theorem 2). With iterative updates of the least-squares estimator, the per step computational complexity is $\mathcal{O}(Ld^2|\mathcal{X}|)$ if the UCB action is computed by a simple enumeration over all actions.

4.1 Regret analysis

Recall that x_t is the action that the UCB algorithm selects at time t , ψ_t the corresponding feature vector in Ψ_t and we define $\psi_t^* = \arg \max_{\psi \in \Psi_t} \psi^\top \theta$. We show that the regret \mathcal{R}_T is bounded in terms of the regret $\mathcal{R}_T^{UCB} := \sum_{t=1}^T \psi_t^{*\top} \theta - \psi_t^\top \theta$ of the UCB algorithm on the contextual bandit defined by the sequence of action feature sets Ψ_t .

Lemma 3. *The regret of Algorithm 1 with the expected feature set Ψ_t is bounded at time T with probability at least $1 - \delta$,*

$$\mathcal{R}_T \leq \mathcal{R}_T^{UCB} + 4\sqrt{2T \log \frac{1}{\delta}}.$$

Further, if the algorithm uses the sample based features $\tilde{\Psi}_t$ with $L = t$ at iteration t , the regret is bounded at time T with probability at least $1 - \delta$,

$$\mathcal{R}_T \leq \mathcal{R}_T^{UCB} + 4\sqrt{2T \log \frac{|\mathcal{X}|\pi T}{3\delta}}.$$

Proof. Consider first the case where we use the expected features $\bar{\psi}_{x_t,\mu_t}$. We add and subtract $(\bar{\psi}_{x_t,\mu_t} - \bar{\psi}_{x_t,\mu_t}^*)^\top \theta$ and use $\bar{\psi}_{x_t,\mu_t}^* \leq \psi_t^{*\top} \theta$ to bound the regret by

$$R_T \leq \mathcal{R}_T^{UCB} + \sum_{t=1}^T D_t,$$

where we defined $D_t = (\phi_{x_t^*, c_t} - \bar{\psi}_{x_t^*, \mu_t} + \bar{\psi}_{x_t, \mu_t} - \phi_{x_t, c_t})^\top \theta$. It is easy to verify that $\mathbb{E}_{c_t \sim \mu_t}[D_t | \mathcal{F}_{t-1}, \mu_t, x_t] = 0$, that is D_t is a martingale difference sequence with $|D_t| \leq 4$ and $M_T = \sum_{t=1}^T D_t$ is a martingale. The first part of the lemma therefore follows from Azuma-Hoeffding's inequality (Lemma 4, Appendix). For the sample-based version, the reasoning is similar, but we need to ensure that the features $\tilde{\psi}_{x, \mu_t} = \frac{1}{L} \sum_{l=1}^L \phi_{x, \tilde{c}_l}$ are sufficiently concentrated around their expected counterparts $\bar{\psi}_{x, \mu_t}$ for any $x \in \mathcal{X}$ and $t \in \mathbb{N}$. We provide details in Appendix A.1. \square

Proof of Theorem 1. Clearly, the lemma gives a regret bound for Algorithm 1 if the regret term \mathcal{R}_T^{UCB} is bounded. This is not immediate, however, as the regret observation $y_t = \phi_{x_t, c_t}^\top \theta + \epsilon_t$ is generated from a different feature vector than the feature $\psi_t \in \Psi_t$ that is chosen by UCB. Note that in general, it is not even true that $\phi_{x_t, c_t} \in \Psi_t$. However, a closer inspection of the reward signal reveals that y_t can be written as

$$y_t = \psi_t^\top \theta + \xi_t + \epsilon \quad \text{with} \quad \xi_t := (\phi_{x_t, c_t} - \psi_t)^\top \theta \quad (2)$$

For the variant that uses the expected features $\psi_t = \bar{\psi}_{x_t, \mu_t}$, our construction already ensures that $\mathbb{E}[\xi_t | \mathcal{F}_{t-1}, \mu_t, x_t] = \mathbb{E}[\phi_{x_t, c_t} - \bar{\psi}_{x_t, \mu_t} | \mathcal{F}_{t-1}, \mu_t, x_t]^\top \theta = 0$. Note that the distribution of ξ_t depends on x_t and is therefore heteroscedastic in general. However, by boundedness, $|\xi_t| \leq 2$ and hence ξ_t is 2-subgaussian, which allows us to continue with a homoscedastic noise bound. We see that y_t acts like an observation of $\psi_t^\top \theta$ perturbed by $\sqrt{4 + \sigma^2}$ -subgaussian noise (for two independent random variables X and Y that are σ_1 - and σ_2 -subgaussian respectively, $X + Y$ is $\sqrt{\sigma_1^2 + \sigma_2^2}$ -subgaussian). Therefore, the construction of the confidence bounds for the least squares estimator w.r.t. ψ_t remains valid at the cost of an increased variance proxy, and we are required to use β_t with $\rho = \sqrt{4 + \sigma^2}$ in the definition of the confidence set. The regret bound for the UCB algorithm (Lemma 2) and an application of the union bound completes the proof for this case. When we use the sample-based features $\tilde{\psi}_{x, \mu_t}$, the noise term ξ_t can be biased, because x_t depends on the sampled features and $\mathbb{E}[\tilde{\psi}_{x_t, \mu_t} | \mathcal{F}_{t-1}, \mu_t] \neq \bar{\psi}_{x_t, \mu_t}$. This bias carries on to the least-squares estimator, but can be controlled by a more careful analysis. See Appendix A.2 for details. \square

4.2 When the context realization is observed

We now turn our attention to the alternative setting, where it is possible to observe the realized context c_t (e.g. actual weather measurements) after the learner has chosen x_t . In Algorithm 1, so far our estimate $\hat{\theta}_t$ only uses the data $\{(x_s, \mu_s, y_s)\}_{s=1}^t$, but with the context observation we have $\{(x_s, c_s, y_s)\}_{s=1}^t$ available. It makes sense to use the additional information to improve our estimate $\hat{\theta}_t$, and as we show below this reduces the amount the UCB algorithm explores. The pseudo code of the modified algorithm is given in Algorithm 2 (Appendix B), where the only difference is that we replaced the estimate of θ by the least squares estimate $\hat{\theta}_t = \arg \min_{\theta' \in \mathbb{R}^d} \sum_{s=1}^t (\phi_{x_s, c_s}^\top \theta' - y_s)^2 + \lambda \|\theta'\|_2^2$. Since now the observation noise $\epsilon_t = y_t - \phi_{x_t, c_t}^\top \theta$ is only σ -subgaussian (instead of $\sqrt{4 + \sigma^2}$ -subgaussian), we can use the smaller scaling factor β_t with $\rho = \sigma$ to obtain a tighter upper confidence bound.

Theorem 2. *The regret of Algorithm 2 based on the expected feature sets Ψ_t and $\beta_t = \beta_t(\sigma, \delta/3)$ is bounded with probability at least $1 - \delta$ by*

$$\mathcal{R}_T \leq \beta_T \sqrt{8T \log \left(\frac{\det V_T}{\det V_0} \right)} + 4(1 + \lambda^{-1/2} \beta_T) \sqrt{2T \log \frac{3}{\delta}}$$

The importance of this result is that it justifies the use of the smaller scaling β_t of the confidence set, which affects the action choice of the UCB algorithm. In practice, β_t has a large impact on the amount of exploration, and a tighter choice can significantly reduce the regret as we show in our experiments. We note that in this case, the reduction to the regret bound of UCB is slightly more involved than previously. As before, we use Lemma 3 to reduce a regret bound on \mathcal{R}_T to the regret \mathcal{R}_T^{UCB} that the UCB algorithm obtains on the sequence of context-feature sets Ψ_t . Since now, the UCB action is based on tighter confidence bounds, we expect the regret \mathcal{R}_T^{UCB} to be smaller, too. This does not follow directly from the UCB analysis, as there the estimator is based on the features $\bar{\psi}_{x_t, \mu_t}$ instead of ϕ_{x_t, c_t} . We defer the complete proof to Appendix B.1. There we also show a similar result for the sample based feature sets $\tilde{\Psi}_t$ analogous to Theorem 1.

4.3 Kernelized stochastic bandits with context distributions

In the kernelized setting, the reward function $f : \mathcal{X} \times \mathcal{C} \rightarrow \mathbb{R}$ is a member of a known reproducing kernel Hilbert space (RKHS) \mathcal{H} with kernel function $k : (\mathcal{X} \times \mathcal{C})^2 \rightarrow \mathbb{R}$. In the following, let $\|\cdot\|_{\mathcal{H}}$ be the Hilbert norm and we denote by $k_{x,c} := k(x, c, \cdot, \cdot) \in \mathcal{H}$ the kernel features. For the analysis we further make the standard boundedness assumption $\|f\| \leq 1$ and $\|k_{x,c}\| \leq 1$. We provide details on how to estimate f given data $\{(x_s, \mu_s, y_s)\}_{s=1}^t$ with uncertain context. As in the linear case, the estimator \hat{f}_t can be defined as an empirical risk minimizer with parameter $\lambda > 0$,

$$\hat{f}_t = \arg \min_{f \in \mathcal{H}} \sum_{s=1}^t (\mathbb{E}_{c \sim \mu_t} [f(x_s, c)] - y_s)^2 + \lambda \|f\|_{\mathcal{H}}^2. \quad (3)$$

In the literature this is known as *distributional risk minimization* (Muandet et al., 2017, Section 3.7.3). The following representer theorem shows, that the solution can be expressed as a linear combination of kernel mean embeddings $\bar{k}_{x,\mu} := \mathbb{E}_{c \sim \mu} [k_{x,c}] \in \mathcal{H}$.

Theorem 3 (Muandet et al. (2012, Theorem 1)). *Any $f \in \mathcal{H}$ that minimizes the regularized risk functional (3) admits a representation of the form $f = \sum_{s=1}^t \alpha_s \bar{k}_{x_s, \mu_s}$ for some $\alpha_s \in \mathbb{R}$.*

It is easy to verify that the solution to (3) can be written as

$$\hat{f}_t(x, c) = k_t(x, c)^\top (K_t + \lambda \mathbf{I})^{-1} y_t \quad (4)$$

where $k_t(x, c) = [\bar{k}_{x_1, \mu_1}(x, c), \dots, \bar{k}_{x_t, \mu_t}(x, c)]^\top$, $(K_t)_{a,b} = \mathbb{E}_{c \sim \mu_b} [\bar{k}_{x_a, \mu_a}(x_b, c)]$ for $1 \leq a, b \leq t$ is the kernel matrix and $y_t = [y_1, \dots, y_t]^\top$ denotes the vector of observations. Likewise, the estimator can be computed from sample based kernel mean embeddings $\tilde{k}_{x,\mu}^l := \frac{1}{l} \sum_{i=1}^l k(x, \tilde{c}_i, \cdot, \cdot) \in \mathcal{H}$ for i.i.d. samples $\tilde{c}_i \sim \mu$. This allows for an efficient implementation also in the kernelized setting, at the usual cost of inverting the kernel matrix. With iterative updates the overall cost amount to $\mathcal{O}(lT^3)$. The cubic scaling in T can be avoided with finite dimensional feature approximations or inducing points methods, e.g. Rahimi and Recht (2008); Mutny and Krause (2018).

The UCB algorithm can be defined using an analogous concentration result for the RKHS setting (Abbasi-Yadkori, 2012). We provide details and the complete kernelized algorithm (Algorithm 3) in Appendix C. The corresponding regret bound is summarized in the following theorem.

Theorem 4. *At any time $T \in \mathbb{N}$, the regret of Algorithm 3 with exact kernel mean embeddings $\bar{k}_{x,c}$ and β_t as defined in Lemma 5 in Appendix C, is bounded with probability at least $1 - \delta$ by*

$$\mathcal{R}_T \leq \beta_T \sqrt{8T \log(\det(\mathbf{I} + (\lambda\rho)^{-1} K_T))} + 4\sqrt{2T \log \frac{2}{\delta}}$$

Also here, the data dependent log-determinant in the regret bound can be replaced with kernel specific bounds, referred to as *maximum information gain* γ_T (Srinivas et al., 2010).

5 Experiments

We evaluate the proposed method on a synthetic example as well as on two benchmarks that we construct from real-world data. Our focus is on understanding the effect of the sample size l used to define the context set $\tilde{\Psi}_t^l$. We compare three different observational modes, with decreasing amount of information available to the learner. First, in the *exact* setting, we allow the algorithm to observe the context realization before choosing an action, akin to the usual contextual bandit setting. Note that this variant possibly obtains negative reward on the regret objective (1), because x_t^* is computed to maximize the expected reward over the context distribution independent of c_t . Second, in the *observed* setting, decisions are based on the context distribution, but the regression uses the exact context realization. Last, only the context distribution is used for the *hidden* setting. We evaluate the effect of the sample sizes $l = 10, 100$ and compare to the variant that uses that exact expectation of the features. As common practice, we treat the confidence parameter β_T as tuning parameter that we choose to minimize the regret after $T = 1000$ steps. For the final evaluation, we show averaged performance over 100 trials. Below we provide details on the experimental setup and the evaluation is shown in Figure 1.

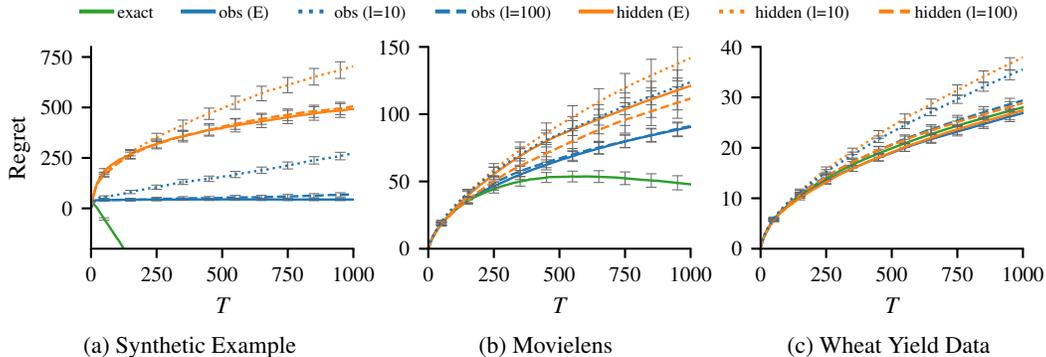


Figure 1: The plots show cumulative regret as defined in (1). As expected, the variant that does not observe the context realization (*hidden*) is out-performed by the variant that uses the context realization for regression (*obs*) but not for choosing the actions. The sample size l used to construct the feature sets from the context distribution has a significant effect on the regret, where with $l = 100$ performance is already competitive with the policy that uses the exact expectation over features (E). The *exact* baseline, which has access to the context realization before taking the decision, quickly obtains negative regret on the benchmark (a) and (b), as the regret objective (1) compares to the action maximizing the expected reward. Error bars show two times standard error over 100 trials.

Synthetic Example As a simple synthetic benchmark we set the reward function to $f(x, c) = \sum_{i=1}^5 (x_i - c_i)^2$, where both actions and context are vectors in \mathbb{R}^5 . We choose this quadratic form to create a setting where the optimal action strongly depends on the context c_i . As linear parametrization we choose $\phi(x, c) = (x_1^2, \dots, x_5^2, c_1^2, \dots, c_5^2, x_1 c_1, \dots, x_5 c_5)$. The action set consists of $k = 100$ elements that we sample at the beginning of each trail from a standard Gaussian distribution. For the context distribution, we first sample a random element $m_t \in \mathbb{R}^5$, again from a multivariate normal distribution, and then set $\mu_t = \mathcal{N}(m_t, \mathbf{1})$. Observation noise is Gaussian with standard deviation 0.1.

MovieLens Data Using matrix factorization we construct 6-dimensional features for user ratings of movies provided in the *movielens-1m* dataset (Harper and Konstan, 2016). We use the learned embedding as ground truth to provide responses to the algorithm. All ratings that we generate in the simulated environment, are rounded to half-integers between 0 and 5 likewise the actual ratings. Therefore our model is miss-specified in this experiment. Besides the movie ratings, the data set provides basic demographic data for each user. In the interactive setting, the context realization is a randomly sampled user from the data. The context distribution is set to the empirical distribution of users in the dataset with the same demographic data. The setup is motivated by a setting where the system interacts with new users, for which we already obtained the basic demographic data, but not yet the exact user’s features (that in collaborative filtering are computed from the user’s ratings). We provide further details in Appendix D.1.

Crop Yield Data We use a wheat yield dataset that was systematically collected by the Agroscope institute in Switzerland over 15 years on 10 different sites. For each site and year, additionally a 16-dimensional suitability factor based on recorded weather conditions is available. The dataset contains 8849 yield measurements for 198 crops. From this we construct a data set $\mathcal{D} = \{(x_i, w_i, y_i)\}$ where x_i is the identifier of the tested crop, $w_i \in \mathbb{R}^{16+10}$ is a 16 dimensional suitability factor obtained from weather measurements augmented with a 1-hot encoding for each site, and y_i is the normalized crop yield. We fit a bilinear model, such that $y_i \approx w_i^T W V_{x_i}$ to obtain 5-dimensional features V_x for each variety x and site features $w^T W$ that take the weather conditions w of the site into account. Again we use the model to provide responses in the interactive setting. For the bandit setting, our goal is to provide a crop recommendation to maximize yield on a given site with characteristics w . Since w is based on weather measurements that are not available ahead of time, we set the context distribution such that each feature of w is perturbed by a Gaussian distribution centered around the true w . The variance of the perturbation we obtain from the empirical variance of the features for the current site over all 15 years. Further details can be found in Appendix D.2.

6 Related Work

There is a large array of work on bandit algorithms, for a survey see Bubeck and Cesa-Bianchi (2012) or the book by Lattimore and Szepesvári (2018). Of specific interest for us is the *stochastic contextual bandit problem*, where the learner chooses actions after seeing a context; and the goal is to compete with a class of policies, that map contexts to actions. This is akin to reinforcement learning (Sutton and Barto, 2018), where a policy maps states to actions; but the contextual bandit problem is different in that the sequence of contexts is typically allowed to be *arbitrary* (even adversarially chosen), and does not necessarily follow a specific transition model. The contextual bandit problem in this formulation dates back to at least Abe and Long (1999) and Langford and Zhang (2007). The perhaps best understood instance of this model is the *linear contextual bandit*, where the reward function is a linear map of feature vectors (Auer et al., 2002). One of the most popular algorithms is the Upper Confidence Bound (UCB) algorithm, first introduced by Auer (2002) for the multi-armed bandit problem, and later extended to the linear case by Li et al. (2010). Analysis of this algorithm was improved by Dani et al. (2008), Abbasi-Yadkori et al. (2011) and Li et al. (2019), where the main technical challenge is to construct tight confidence sets for an online version of the least squares estimator. Alternative exploration strategies have been considered as well, for instance Thompson sampling (Thompson, 1933), which was analyzed for the linear model by Agrawal and Goyal (2013) and Abeille and Lazaric (2017). This policy can be computationally more efficient than UCB, but the best known frequentist regret bounds are worse in the dependency of the dimension by a factor \sqrt{d} . Other notable approaches include an algorithm that uses a perturbed data history as exploration mechanism (Kveton et al., 2019), or a *mostly greedy* algorithm that leverages the randomness in the context to obtain sufficient exploration (Bastani et al., 2017). A generalization of the linear setting are *kernelized bandits*, where the reward function is a member of a given reproducing kernel Hilbert space (RKHS). This setting is closely related to Bayesian optimization (Mockus, 1982). Again, the analysis hinges on the construction of confidence sets as well as bounding a quantity referred to as *information gain* by the decay of the kernel’s eigenspectrum. An analysis of the UCB algorithm in this setting was provided by Srinivas et al. (2010) and later refined by Abbasi-Yadkori (2012); Valko et al. (2013); Chowdhury and Gopalan (2017); Durand et al. (2018). Thompson sampling was analyzed in the RKHS setting by Chowdhury and Gopalan (2017) and UCB for the contextual variant was introduced by Krause and Ong (2011). Interestingly, our analysis shows a reduction where we construct an instance of the linear bandit problem, where the distribution of the observation noise depends on the action, also referred to as *heteroscedastic* bandits. Heteroscedastic bandits with linear and RKHS reward function where previously considered by Hsieh et al. (2018) and Kirschner and Krause (2018). The perhaps closest work to ours is by Yun et al. (2017), who introduce a linear contextual bandit model where the observed feature is perturbed by noise and the objective is to compete with the best policy that has access to the unperturbed feature vector. The main difference to our setting is that we assume that the environment provides a distribution of feature vectors (instead of a single, perturbed vector) and we compute the best action as a function of the distribution. As a consequence, we are able to obtain $\mathcal{O}(\sqrt{T})$ regret bounds without further assumptions on the context distribution, while Yun et al. (2017) get $\mathcal{O}(T^{7/8})$ with identical noise on each feature, and $\mathcal{O}(T^{2/3})$ for Gaussian feature distributions. In principle, it is also possible to apply adversarial contextual bandit algorithms in our setting, for example the EXP4 algorithm of Auer et al. (2002) or ILTCB of Agarwal et al. (2014). Here, the objective is to compete with the best policy in a given class of policies, which in our setting would require to work with a covering of the set of distributions $\mathcal{P}(\mathcal{C})$. However, these algorithms do not exploit the linear reward assumption and, arguably, are less practical for our setting.

7 Conclusion

We introduced *context distributions* for stochastic bandits, a model that is naturally motivated in many applications and allows to capture the learner’s uncertainty in the context realization. The method we propose is based on the UCB algorithm, and in fact, both our model and algorithm strictly generalize the standard setting in the sense that we recover the usual model and the UCB algorithm if the environment chooses only Dirac delta distributions. The most practical variant of the proposed algorithm requires only sample access to the context distributions and satisfies a high-probability regret bound that is order optimal in the feature dimension and the horizon up to logarithmic factors.

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A Proof Details

Lemma 4 (Azuma-Hoeffdings). *Let M_i be a martingale on a filtration \mathcal{F}_i with almost surely bounded increments $|M_i - M_{i-1}| < B$. Then*

$$\mathbb{P}[M_n - M_0 > s] \leq \exp\left(-\frac{s^2}{2nB^2}\right)$$

A.1 Proof of Lemma 3

To bound the regret of Algorithm 1 with sampled features $\tilde{\psi}_{x,\mu_t}$, this time we add and subtract $(\tilde{\psi}_{x_t^*,\mu_t} + \bar{\psi}_{x_t,\mu_t} + \tilde{\psi}_{x_t,\mu_t})^\top \theta$ to the regret to find

$$\mathcal{R}_T \leq \mathcal{R}_T^{UCB} + 4\sqrt{2T \log \frac{1}{\delta}} + \sum_{t=1}^T (\tilde{\psi}_{x_t,\mu_t} - \bar{\psi}_{x_t,\mu_t})^\top \theta.$$

By the same reasoning as for the expected features, we bounded

$$\sum_{t=1}^T (\phi_{x_t^*,c_t} - \tilde{\psi}_{x_t^*,\mu_t} + \bar{\psi}_{x_t,\mu_t} - \phi_{x_t,c_t})^\top \theta \leq 4\sqrt{2T \log \frac{1}{\delta}} \quad (5)$$

using Azuma-Hoeffding's inequality. We are left with a sum $\sum_{t=1}^T (\tilde{\psi}_{x_t,\mu_t} - \bar{\psi}_{x_t,\mu_t})^\top \theta$ that is more intricate to bound because x_t depends on the samples $\tilde{c}_{t,l}$ that define $\tilde{\psi}_{x_t,\mu_t}$. Instead, first consider a fixed $x \in \mathcal{X}$. Then, by $(\tilde{\psi}_{x,\mu_t} - \bar{\psi}_{x,\mu_t})^\top \theta \leq 2$ and Azuma-Hoeffding's inequality with probability at least $1 - \delta$,

$$(\tilde{\psi}_{x,\mu_t} - \bar{\psi}_{x,\mu_t})^\top \theta \leq \sqrt{\frac{8}{L} \log \frac{1}{\delta}}. \quad (6)$$

We enforce this to hold for any $x \in \mathcal{X}$ and any time $t \in \mathbb{N}$ by replacing δ by $\frac{6\delta}{|\mathcal{X}|\pi^2 t^2}$ and taking the union bound over the event where (6) holds. By our choice $L = t$ and $\sum_{t=1}^T \frac{1}{\sqrt{t}} \leq 2\sqrt{T}$, we have with probability at least $1 - \delta$, for any $T \in \mathbb{N}$,

$$\sum_{t=1}^T (\tilde{\psi}_{x_t,\mu_t} - \bar{\psi}_{x_t,\mu_t})^\top \theta \leq 4\sqrt{T \log \frac{|\mathcal{X}|\pi^2 T^2}{6\delta}}. \quad (7)$$

A final application of the union bound over the events such that (5) and (7) simultaneously hold, gives

$$R_T \leq \mathcal{R}_T^{UCB} + 4\sqrt{2T \log \frac{|\mathcal{X}|\pi T}{3\delta}}.$$

A.2 Proof of Theorem 1

To bound the regret term \mathcal{R}_T^{UCB} for the case where we uses sample-based feature vectors, the main task is to show a high-probability bound on $\|\theta - \hat{\theta}_t\|_{V_t}$ with observations $y_t = \tilde{\psi}_{x,t}^\top \theta + \xi_t + \epsilon_t$. Recall that $\xi_t = (\phi_{x,c_t} - \tilde{\psi}_{x_t,\mu_t})^\top \theta$, but now we have in general $\mathbb{E}[\xi_t | \mathcal{F}_{t-1}, \mu_t, x_t] \neq 0$, because x_t depends on the sampled features $\tilde{\psi}_{x,\mu_t}$. Basic linear algebra and the triangle inequality show that

$$\|\theta - \hat{\theta}_t\|_{V_t} \leq \left\| \sum_{t=1}^T \tilde{\psi}_{x,t} (\xi_t + \epsilon_t) \right\|_{V_t^{-1}} + \lambda^{1/2} \|\theta\|_2 \quad (8)$$

Recall that we assume $\|\theta\|_2 \leq 1$ to bound the last term. We split the noise process into an unbiased part, that can be controlled by standard results, and the bias that we aim to control by the number of samples L .

$$\left\| \sum_{t=1}^T \tilde{\psi}_{x,t} (\xi_t + \epsilon_t) \right\|_{V_t^{-1}} \leq \left\| \sum_{t=1}^T \tilde{\psi}_{x,t} ((\phi_{x,c_t} - \bar{\psi}_{x_t,\mu_t})^\top \theta + \epsilon_t) \right\|_{V_t^{-1}} + \left\| \sum_{t=1}^T \tilde{\psi}_{x,t} b_t \right\|_{V_t^{-1}} \quad (9)$$

where we defined $b_t = (\bar{\psi}_{x_t, \mu_t} - \tilde{\psi}_{x_t, \mu_t})^\top \theta$. As before, $\mathbb{E}[(\phi_{x_t, c_t} - \bar{\psi}_{x_t, \mu_t})^\top \theta + \epsilon_t | \mathcal{F}_{t-1}, \mu_t, x_t] = 0$, so the first sum can be upper bounded, e.g. by (Abbasi-Yadkori et al., 2011, Theorem 1), to obtain

$$\left\| \sum_{t=1}^T \tilde{\psi}_{x_t, t} ((\phi_{x_t, c_t} - \bar{\psi}_{x_t, \mu_t})^\top \theta + \epsilon_t) \right\|_{V_t^{-1}}^2 \leq 2(4 + \sigma^2) \log \left(\frac{\det(V_t)^{1/2}}{\delta \det(V_0)^{1/2}} \right) \quad (10)$$

To control the sum over bias terms b_t , note that by our choice $L = t$, similar to (7), with probability at least $1 - \delta$, for all $t \in \mathbb{N}$ and $x \in \mathcal{X}$,

$$|b_t| \leq \sqrt{\frac{8}{t} \log \frac{\pi^2 t^2 |\mathcal{X}|}{6\delta}}$$

Now set $b = [b_1, \dots, b_t]$ and $A = [\tilde{\psi}_{x_1, 1}, \dots, \tilde{\psi}_{x_t, t}]^\top \in \mathbb{R}^{d \times t}$. The matrix inequality

$$A^\top (AA^\top + \lambda \mathbf{I}_d)^{-1} A \leq \mathbf{I}_t \quad (11)$$

follows by using a SVD decomposition. This implies

$$\left\| \sum_{t=1}^T \tilde{\psi}_{x_t, t} b_t \right\|_{V_t^{-1}}^2 = \|Ab\|_{(AA^\top + \lambda \mathbf{I}_d)^{-1}}^2 \leq \|b\|_2^2 \leq 8 \log(T) \log \left(\frac{\pi^2 T^2 |\mathcal{X}|}{6\delta} \right)$$

The last step follows from $\sum_{s=1}^t \frac{1}{s} \leq \log(t)$. Going back to (8), we have now bounded all the terms, so with a final union bound over (10) and (12) we obtain with probability at least $1 - \delta$,

$$\begin{aligned} \|\theta - \hat{\theta}_t\|_{V_t} &\leq \left\| \sum_{t=1}^T \tilde{\psi}_{x_t, t} ((\phi_{x_t, c_t} - \bar{\psi}_{x_t, \mu_t})^\top \theta + \epsilon_t) \right\|_{V_t^{-1}} + \left\| \sum_{t=1}^T \tilde{\psi}_{x_t, t} b_t \right\|_{V_t^{-1}} + \lambda \|\theta\|_2 \\ &\leq \sqrt{2(4 + \sigma^2) \log \left(\frac{2 \det(V_t)^{1/2}}{\delta \det(V_0)^{1/2}} \right)} + \sqrt{8 \log(T) \log \left(\frac{\pi^2 T^2 |\mathcal{X}|}{3\delta} \right)} + \lambda. \end{aligned} \quad (12)$$

Finally, we invoke Lemma 2 with $\beta_t = \tilde{\beta}_t$, where

$$\tilde{\beta}_t := \sqrt{2(4 + \sigma^2) \log \left(\frac{2 \det(V_t)^{1/2}}{\delta \det(V_0)^{1/2}} \right)} + \sqrt{8 \log(T) \log \left(\frac{\pi^2 T^2 |\mathcal{X}|}{3\delta} \right)} + \lambda \quad (13)$$

to obtain a bound on \mathcal{R}_T^{UCB} ,

$$\mathcal{R}_T^{UCB} \leq \tilde{\beta}_T \sqrt{8T \log \left(\frac{\det V_T}{\det V_0} \right)}. \quad (14)$$

This concludes the proof.

B UCB with Context Distributions and Observed Context

B.1 Proof of Theorem 2

From Lemma 3 we obtain with probability at least $1 - \delta$,

$$\mathcal{R}_T \leq \sum_{t=1}^T \psi_t^{*\top} \theta - \bar{\psi}_{x_t, \mu_t}^\top \theta + 4 \sqrt{2T \log \frac{1}{\delta}} \quad (15)$$

What is different now, is that the sum $\sum_{t=1}^T \psi_t^{*\top} \theta - \bar{\psi}_{x_t, \mu_t}^\top \theta$ contains actions x_t , that are computed from a more precise estimator $\hat{\theta}_t$, hence we expect the regret to be smaller. This does not follow directly from UCB analysis, as there the estimator is computed with the features $\bar{\psi}_{x_t, \mu_t}$.

We start with the usual regret analysis, making use of the confidence bounds (Lemma 1) and the definition of the UCB action. Denote $\psi_t := \psi_{x_t, \mu_t}$ in the following.

$$\sum_{t=1}^T \psi_t^{*\top} \theta - \psi_t^\top \theta \leq \sum_{t=1}^T \psi_t^{*\top} \hat{\theta}_t + \beta_t \|\psi_t^*\|_{V_t^{-1}} - (\phi_t^\top \hat{\theta}_t - \beta_t \|\psi_t\|_{V_t^{-1}}) \leq 2\beta_T \sum_{t=1}^T \|\psi_t\|_{V_t^{-1}} \quad (16)$$

Algorithm 2 UCB for linear stochastic bandits with context distributions and observed context

Initialize $\hat{\theta} = 0 \in \mathbb{R}^d$, $V_0 = \lambda \mathbf{I} \in \mathbb{R}^{d \times d}$

For step $t = 1, 2, \dots, T$:

Environment chooses $\mu_t \in \mathcal{P}(\mathcal{C})$ // context distribution
Learner observes μ_t

Set $\Psi_t = \{\psi_{x,\mu_t} : x \in \mathcal{X}\}$ with $\bar{\psi}_{x,\mu_t} = \mathbb{E}_{\mu_t}[\phi_{x,c}]$ // expected version
Alternatively, sample $c_{t,1}, \dots, c_{t,L}$ for $L = t$, // or sampled version

Set $\Psi_t = \{\tilde{\psi}_{x,\mu_t} : x \in \mathcal{X}\}$ with $\tilde{\psi}_{x,\mu_t} := \frac{1}{L} \sum_{i=1}^L \phi_{x,\tilde{c}_i}$

Run UCB step with Ψ_t as context set // reduction

Choose action $x_t = \arg \max_{x \in \mathcal{X}} \psi_{x,\mu_t}^\top \hat{\theta}_{t-1} + \beta_t(\sigma) \|\psi_{x,\mu_t}\|_{V_{t-1}^{-1}}$ // UCB action

Environment samples $c_t \sim \mu_t$
Learner observes $y_t = \phi_{x_t,c_t}^\top \theta + \epsilon_t$ and c_t // reward and context observation

Update $V_t = V_{t-1} + \phi_{x_t,c_t} \phi_{x_t,c_t}^\top$, $\hat{\theta}_t = V_t^{-1} \sum_{s=1}^t \phi_{x_s,c_s} y_s$ // least-squares update

From here, the standard analysis proceeds by using Cauchy-Schwarz to obtain an upper bound on $\sum_{t=1}^T \|\psi_t\|_{V_t^{-1}}$ and then the argument proceeds by simplifying the sum $\sum_{t=1}^T \|\psi_t\|_{V_t^{-1}}$. The simplification doesn't work here because V_t is defined on the realized features ϕ_{x_t,c_t} and not ψ_t . Instead we require the following intermezzo.

$$\sum_{t=1}^T \|\psi_t\|_{V_t^{-1}} = \sum_{t=1}^T \|\phi_{x_s,c_s}\|_{V_t^{-1}} + \|\psi_t\|_{V_t^{-1}} - \|\phi_{x_s,c_s}\|_{V_t^{-1}} \leq \sum_{t=1}^T \|\phi_{x_s,c_s}\|_{V_t^{-1}} + \sum_{t=1}^T S_t, \quad (17)$$

where we defined $S_t = \|\psi_t\|_{V_t^{-1}} - \|\phi_{x_s,c_s}\|_{V_t^{-1}}$. We show that $\sum_{t=1}^T S_t$ is a supermartingale. For the expected features $\bar{\psi}_{x,\mu_t} = \mathbb{E}_{c \sim \mu_t}[\phi_{x,c}]$, note that Jensen's inequality yields $\|\mathbb{E}_{c \sim \mu_t}[\phi_{x,c}]\|_{V_t^{-1}} \leq \mathbb{E}_{c \sim \mu_t}[\|\phi_{x,c}\|_{V_t^{-1}}]$ for all $x \in \mathcal{X}$. From this we obtain $\mathbb{E}[S_t | \mathcal{F}_{t-1}, \mu_t] \leq 0$. Finally, note that $\|\phi_{x_s,c_s}\|_{V_t^{-1}} \leq \lambda^{-1/2} \|\phi_{x_s,c_s}\|_2 \leq \lambda^{-1/2}$ and $|S_t| \leq 2\lambda^{-1/2}$, hence by Azuma-Hoeffdings inequality with probability at least $1 - \delta$, $\sum_{t=1}^T S_t \leq 2\lambda^{-1/2} \sqrt{2T \log \frac{1}{\delta}}$.

From here we complete the regret analysis by bounding $\sum_{t=1}^T \|\phi_{x_s,c_s}\|_{V_t^{-1}}$ with the standard argument. Write $\phi_t := \phi_{x_t,c_t}$. First using Cauchy-Schwarz and then $\|\phi_t\|_2 \leq 1$ as well as $u \leq 2 \log(1+u)$ for $u \leq 1$, it follows that

$$\sum_{t=1}^T \|\phi_t\|_{V_t^{-1}} \leq \sqrt{T \sum_{t=1}^T \|\phi_t\|_{V_t^{-1}}^2} \leq \sqrt{2T \sum_{t=1}^T \log(1 + \|\phi_t\|_{V_t^{-1}}^2)} = \sqrt{2T \log \left(\frac{\det V_T}{\det V_0} \right)} \quad (18)$$

The last equality essentially follows from an application of the Sherman-Morrison formula on the matrix $V_t = \sum_{s=1}^t \phi_{x_s,c_s} \phi_{x_s,c_s}^\top + \lambda \mathbf{I}_d$, compare e.g. (Abbasi-Yadkori et al., 2011, Lemma 11). It remains to assemble the results from equations (16)-(18), and a final application of the union bound completes the proof.

With a bit of extra work, one can obtain a similar bound for the sample-based features. Using Jensen's inequality we get $\|\frac{1}{L} \sum_{i=1}^L \phi_{x,\tilde{c}_i}\|_{V_t^{-1}} \leq \frac{1}{L} \sum_{i=1}^L \|\phi_{x,\tilde{c}_i}\|_{V_t^{-1}}$, but now the action x_t depends on the samples $\tilde{c}_{t,i}$ that define the features $\tilde{\psi}_{x_t,\mu_t}$ and the previous direct argument does not work. The strategy around is the same as in the proof of Lemma 3. For fixed $x \in \mathcal{X}$, $\frac{1}{L} \sum_{i=1}^L \|\phi_{x,\tilde{c}_i}\|_{V_t^{-1}}$ concentrates around $\mathbb{E}_{c \sim \mu_t}[\|\phi_{x,c}\|_{V_t^{-1}}]$ at a rate $1/\sqrt{L}$, fast enough to get $\mathcal{O}(\sqrt{T})$ regret if we set $L = t$ in iteration t . A careful application of the union bound (again over all $x \in \mathcal{X}$) completes the proof.

C Kernelized UCB with Context Distributions

C.1 Proof of Theorem 4

To bound the regret we proceed like in the linear case. First, define $\bar{f}_t(x) = \mathbb{E}_{\mu_t}[f(x, c)|\mathcal{F}_{t-1}]$. Analogously to Lemma 3, an application of Azuma-Hoeffding's inequality yields with probability at least $1 - \delta$.

$$\mathcal{R}_T \leq \sum_{t=1}^T \bar{f}_t(x_t^*) - \bar{f}_t(x_t^*) + 4\sqrt{2T \log \frac{1}{\delta}} \quad (19)$$

To bound the sum, we need to understand the concentration behavior of $|\hat{f}_t(x) - \bar{f}_t(x_t^*)|$, where we denote $\hat{f}_t(x) = \mathbb{E}_{\mu_t}[\hat{f}_t(x, c)|\mathcal{F}_{t-1}, \mu_t]$. Recall that

$$\hat{f}_t(x, c) = k_t(x, c)^\top (K_t + \lambda \mathbf{I})^{-1} y_t \quad (20)$$

where $k_t(x, c) = [\bar{k}_{x_1, \mu_1}(x, c), \dots, \bar{k}_{x_t, \mu_t}(x, c)]^\top$, $(K_t)_{i,j} = \mathbb{E}_{c' \sim \mu_j}[\bar{k}_{x_i, \mu_i}(x_j, c')]$ for $1 \leq i, j, \leq t$ is the kernel matrix and $y_t = [y_1, \dots, y_t]^\top$ denotes the vector of observations. Define further $\bar{k}_t(x) = \mathbb{E}_{c \sim \mu_t}[k_t(x, c)]$.

Note that we can compute $\hat{f}_t(x) = \langle \hat{f}, k_{x, \mu_t} \rangle = \bar{k}_t(x)^\top (K_t + \lambda \mathbf{I})^{-1} y_t$ according to the inner product $\langle \cdot, \cdot \rangle$ on \mathcal{H} . Concentration bounds for the kernel least squares estimator, that hold for adaptively collected data, are well understood by now, see Srinivas et al. (2010); Abbasi-Yadkori (2012); Chowdhury and Gopalan (2017); Durand et al. (2018). For instance, as a direct corollary of (Abbasi-Yadkori, 2012, Theorem 3.11), we obtain

Lemma 5. *For any stochastic sequence $\{(x_t, \mu_t, y_t)\}_{t \in \mathbb{N}}$, where $y_t = f(x_t, c_t) + \epsilon_t$ with σ -subgaussian noise ϵ_t and $c_t \sim \mu_t$, the kernel-least squares estimate (3) satisfies with probability at least $1 - \delta$, at any time t and for any $x \in \mathcal{X}$,*

$$|\hat{f}_t(x) - \bar{f}_t(x_t^*)| \leq \beta_t \sigma_t(x).$$

Here we denote,

$$\beta_t = \rho \left(\sqrt{2 \log \left(\frac{\det(\mathbf{I} + (\lambda \rho)^{-1} K_t)^{1/2}}{\delta} \right)} + \lambda^{1/2} \|f\|_{\mathcal{H}} \right),$$

$$\sigma_t^2(x) = \frac{1}{\lambda} (\langle k_{x, \mu_t}, k_{x, \mu_t} \rangle - k_t(x)^\top (K_t + \lambda \mathbf{I})^{-1} k_t(x)),$$

and $\rho = \sqrt{4 + \sigma^2}$ is the subgaussian variance proxy of the observation noise $\rho_t = y_t - \bar{f}_t(x)$.

Using the confidence bounds, we define the UCB action,

$$x_t = \arg \max_{x \in \mathcal{X}} \hat{f}_t(x) + \beta_t \sigma_t(x). \quad (21)$$

It remains to bound the regret of the UCB algorithm. The proof is standard except that we use Lemma 5 to show concentration of the estimator. For details see (Abbasi-Yadkori, 2012, Theorem 4.1) or (Chowdhury and Gopalan, 2017, Theorem 3).

D Details on the Experiments

We tune β_t over the values $\{0.5, 1, 2, 5, 10\}$. In the synthetic experiment we set $\beta_t = 2$ for the *exact* and *observed* variant and $\beta_t = 10$ for the *hidden* experiment. In the both experiments based on real-world data, we set $\beta_t = 1$ for all variants.

D.1 Movielens

We use matrix factorization based on singular value decomposition (SVD) (Koren et al., 2009) which is implemented in the Surprise library (Hug, 2017) to learn 6 dimensional features v_u, w_m for users u and movies m (this model obtains a RMSE ≈ 0.88 over a 5-fold cross-validation). In the linear parameterization this corresponds to 36-dimensional feature $\phi_{m,u} = v_u w_m^\top$. We use the demographic data (gender, age and occupation) to group the users. The realized context is set to a random user in the data set, and the context distribution is defined as the empirical distribution over users within the same group as the chosen user.

Algorithm 3 UCB for RKHS bandits with context distributions

Initialize \hat{f}_0, σ_0
For step $t = 1, 2, \dots, T$:
 Environment chooses $\mu_t \in \mathcal{P}(\mathcal{C})$ // context distribution
 Learner observes μ_t

 // definitions for expected version
 $[k_t(x)]_s := \mathbb{E}_{c \sim \mu_t} [k_{x_s, \mu_s}(x, c)]$ for $s = 1, \dots, t-1$
 $s_t(x) := \mathbb{E}_{c \sim \mu_t, c' \sim \mu_t} [k(x, c, x, c')]$

 // definitions for sampled version
 Sample $\tilde{c}_{t,1}, \dots, \tilde{c}_{t,L} \sim \mu_t$ // sample context distribution
 $[k_t(x)]_s := \frac{1}{L} \sum_{i=1}^L k_{x_s, \mu_s}(x, \tilde{c}_i)$ for $s = 1, \dots, t-1$
 $s_t(x) := \frac{1}{L^2} \sum_{i,j=1}^L k(x, \tilde{c}_{t,i}, x, \tilde{c}_{t,j})$

 $\hat{f}_t(x) := k_t(x)^\top (K_t + \lambda \mathbf{I})^{-1} y_t$ // compute estimate
 $\sigma_t^2(x) := \frac{1}{\lambda} (s_t(x) - k_t(x)^\top (K_t + \lambda \mathbf{I})^{-1} k_t(x))$ // confidence width

 Set β_t as in Lemma 5
 Choose action $x_t \in \arg \max_{x \in \mathcal{X}} \hat{f}_{t-1}(x) + \beta_t \sigma_{t-1}(x)$ // UCB action

 Environment provides $y_t = f(x_t, c_t) + \epsilon$ where $c_t \sim \mu_t$ // reward observation
 Store $y_{t+1} := [y_1, \dots, y_t]$ // observation vector

 // Kernel matrix update, expected version
 $K_{t+1}(x) := \langle [k_{x_a, \mu_a}, k_{x_b, \mu_b}]_{1 \leq a, b \leq t} \rangle$ with $\langle k_{x_a, \mu_a}, k_{x_b, \mu_b} \rangle = \mathbb{E}_{\mu_a, \mu_b} [k(x_a, c_a, x_b, c_b)]$
 $k_{x_t, \mu_t} := \mathbb{E}_{c \sim \mu_t} [k_{x_t, c}]$ // kernel mean embeddings

 // Kernel matrix update, sampled version
 $K_{t+1}(x) := \langle [k_{x_a, \mu_a}, k_{x_b, \mu_b}]_{1 \leq a, b \leq t} \rangle$ with $\langle k_{x_a, \mu_a}, k_{x_b, \mu_b} \rangle = \frac{1}{L^2} \sum_{i,j=1}^L k(x_a, \tilde{c}_{a,i}, x_b, \tilde{c}_{b,j})$
 $k_{x_t, \mu_t} := \frac{1}{L} \sum_{i=1}^L k_{x_t, \tilde{c}_{t,i}} = \frac{1}{L} \sum_{i=1}^L k(x_t, \tilde{c}_{t,i}, \cdot, \cdot)$ // sample kernel mean embeddings

D.2 Crops Yield Dataset

Recall that the data set $\mathcal{D} = \{(x_i, w_i, y_i)\}$ for $i = 1, \dots, 8849$ consists of crop identifiers x_i , normalized yield measurements y_i and site-year features $w_i \in \mathbb{R}^{16+10}$ that are based on 16 suitability factors computed from weather measurements and a 1-hot encoding for each site (out of 10). The suitability factors are based on the work of Holzkämper et al. (2013). To obtain a model for the crop yield responses, we train a bilinear model on the following loss (Koren et al., 2009),

$$\mathcal{L}(W, V) = \sum_{i=1}^n (y_i - w_i^\top W V_{x_i})^2 + \|v_{j_i}\|^2 + \|w_{x_i}^\top W\|_2^2. \quad (22)$$

where $V = (V_{x_j})_{j=1}^k$ are the crop features $V_{x_j} \in \mathbb{R}^5$ and $W \in \mathbb{R}^{26 \times 6}$ is used to compute site features $w_i^\top W$ given the suitability features w_i from the dataset.

We also use an empirical noise function created from the data. Since the data set contains up to three measurements of the same crop on a specific site and a year, we randomly pick a measurement and use its residual w.r.t. the mean of all measurements of the same crop under exactly the same conditions as noise. This way we ensure that the observation noise of our simulated environment is of the same magnitude as the noise on the actual measurements.